Uncertainty Quantification and Data Fusion

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ICSP Tutorial, Bergamo, July 2013 This material is based upon work supported in part by the U.S. Army Research Laboratory and the U.S. Army Research Office under grant

numbers 00101-80683, W911NF-10-1-0246 and W911NF-12-1-0273.

Components of deterministic optimization practice



Components of stochastic optimization practice



- objective and constraint functions
- decision stages, uncertainty model



Components of stochastic optimization practice



- objective and constraint functions
- decision stages, uncertainty model



solutions (with guarantees) recommended decisions

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Example: 100 observations of demand

Possibilities: Use the 100 scenarios; fit to assumed parametric form; Bayesian update; nonparametric estimation

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Example: 100 observations of demand

Possibilities: Use the 100 scenarios; fit to assumed parametric form; Bayesian update; nonparametric estimation



...but with only 5 observations



Almost always soft information available

Use stochastic programming!

Same 5 points, but with epi-splines and soft info. (nonnegativity, continuous differentiable, and decreasing)



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Two roles for stochastic programming



Independent value in data fusion and uncertainty quant.

Engineering, biological, physical systems

- Input: random V ("known" distribution)
- ▶ System function *G*; implicitly defined e.g. by simulation
- Output: random variable

$$X=G(V)$$

Independent value in data fusion and uncertainty quant.

Engineering, biological, physical systems

- Input: random V ("known" distribution)
- ► System function *G*; implicitly defined e.g. by simulation
- Output: random variable

$$X = G(V)$$

Given observations (data) $x^1 = G(v^1), ..., x^{\nu} = G(v^{\nu})$, we seek a description of X:

- mean, standard deviation
- quantile, superquantile
- distribution, density (pdf)

Main challenge: few data points; little relevant data But soft information might be available

Example: M/M/1; 50% of customers delayed for fixed time

X = customer time-in-service; 100 observations Soft info: lsc, $X \ge 0$, pointwise Fisher, unimodal upper tail



Outline

Density estimation as a stochastic program

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- Epi-splines: pliable approximation tools
- Consistency and asymptotics
- Fusion of hard and soft information
- Numerical examples

Formulation of density estimation problem

Kullback-Leibler divergence from density h to density g on R

$$d_{\mathcal{KL}}(h||g) = \int_{-\infty}^{\infty} h(x) \log \frac{h(x)}{g(x)} dx$$

Facts:

 $d_{\mathcal{KL}}(h||g)\geq 0$ for all densities h,g $d_{\mathcal{KL}}(h||g)=0 \Longleftrightarrow h(x)=g(x)$ (Lebesgue) almost every $x\in R$

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Consequently, for density $h^0 \in \mathcal{H}$: If

$$\tilde{h} \in \underset{h}{\operatorname{argmin}} d_{KL}(h^{0}||h)$$

s.t.
$$\int_{-\infty}^{\infty} h(x) dx = 1$$
$$h \ge 0$$
$$h \in \mathcal{H}$$

then

$$\tilde{h} = h^0$$
 a.e.

Let X^0 be a random variable with density h^0

Then,

$$d_{\mathcal{K}L}(h^0||h) = \int_{-\infty}^{\infty} h^0(x) \log \frac{h^0(x)}{h(x)} dx = E\{\log h^0(X^0)\} - E\{\log h(X^0)\}$$

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Then,

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So minimizing $d_{KL}(h^0||h)$ is equivalent to

$$egin{array}{ll} & ilde{h}\in rgmax_{h}E\{\log h(X^{0})\}\ & ext{ s.t. } & \int_{-\infty}^{\infty}h(x)dx &= 1\ & ext{ }h &\geq 0, \ h\in\mathcal{H} \end{array}$$

Of course, expectation is with respect to the true distribution

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Data (sample)
$$X^1, X^2, ..., X^{\nu}$$
 available
 $h^{\nu} \in \operatorname{argmax}_{h} \frac{1}{\nu} \sum_{i=1}^{\nu} \log h(X^i) = \log \left(\prod_{i=1}^{\nu} h(X^i)\right)^{1/\nu}$
s.t. $\int_{-\infty}^{\infty} h(x) dx = 1$
 $h \geq 0, h \in \mathcal{H}$

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 $h \geq 0, h \in \mathcal{H}$

approximates

$$ilde{h} \in \operatorname*{argmax}_{h} E\{ \log h(X^0) \}$$

s.t. $\int_{-\infty}^{\infty} h(x) dx = 1$
 $h \geq 0, h \in \mathcal{H}$

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Approximation is a max log-likelihood problem

Incorporating soft information

$$h^{\nu} \in \operatorname{argmax}_{h} \frac{1}{\nu} \sum_{i=1}^{\nu} \log h(X^{i}) = \log \left(\prod_{i=1}^{\nu} h(X^{i}) \right)^{1/\nu}$$

s.t.
$$\int_{-\infty}^{\infty} h(x) dx = 1$$
$$h \geq 0, \ h \in \mathcal{H}^{\nu} \subset \mathcal{H}$$

where H^{ν} includes essentially any soft information about h^0 :

- support bounds
- density continuity, smoothness
- density shape (unimodal, decreasing, etc.)
- moments
- proximity to known density
- ▶ system knowledge (convex *G*, gradient of *G*)

Outline

Density estimation as a stochastic program

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- Epi-splines: pliable approximation tools
- Consistency and asymptotics
- Fusion of hard and soft information
- Numerical examples

Challenge: infinite-dimensional problems

$$h^{\nu} \in \underset{h}{\operatorname{argmax}} \frac{1}{\nu} \sum_{i=1}^{\nu} \log h(X^{i}) = \log \left(\prod_{i=1}^{\nu} h(X^{i}) \right)^{1/\nu}$$

s.t.
$$\int_{-\infty}^{\infty} h(x) dx = 1$$
$$h \geq 0, h \in H^{\nu} \subset \mathcal{H}$$

Need approximation of $\ensuremath{\mathcal{H}}$ by set of flexible functions given by finite number of parameters

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Exponential epi-spline estimator

Given sample $X^1, ..., X^{\nu}$, the exponential epi-spline estimator of the true density is

$$h^{\nu}=e^{-s^{\nu}},$$

where s^{ν} is an epi-spline



Epi-splines: piecewise polynomial functions



Exponential epi-spline estimator

Given sample $X^1, ..., X^{\nu}$, the exponential epi-spline estimator of the true density is

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where s^{ν} is an epi-spline

Main features:

- finite number of parameters; powerful optimization technology available
- approximates to arbitrary accuracy essentially any function
- easily includes soft information
- substantially more flexible and pliable than 'classical' splines
- nonnegativity achieved automatically

Epi-splines

- Number of partitions N
- Mesh $m = \{m_k\}_{k=0}^N$, where $m_{k-1} < m_k$, k = 1, 2, ..., N

- Estimation on [m₀, m_N]
- Order p

Epi-splines

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- Mesh $m = \{m_k\}_{k=0}^N$, where $m_{k-1} < m_k$, k = 1, 2, ..., N
- Estimation on [m₀, m_N]
- Order p

Definition

e-spl^{*p*}(*m*)= family of (basic) epi-splines of order *p*, with mesh $m = \{m_k\}_{k=0}^N$, consists of:

- functions $s : [m_0, m_N] \rightarrow R$
- ▶ that are polynomials of degree p in each segment (m_{k-1}, m_k), k = 1, 2, ..., N, and
- ► that are finite valued at m₀, m₁, ..., m_N

Representation of epi-spline



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Representation of epi-spline

Every $s \in \operatorname{e-spl}^p(m)$, with $m = \{m_k\}_{k=0}^N$, is uniquely represented by

$$r = (s_0, s_1, ..., s_N, a_1, a_2, ..., a_N), \quad s_k \in R, \ a_k \in R^{p+1},$$

such that

$$s(x) = \langle c(x), r \rangle, \ x \in [m_0, m_N],$$

where

$$c(x) = \begin{cases} (0, ..., 0, 1, x - m_{k-1}, ..., (x - m_{k-1})^p, 0, ..., 0) \\ & \text{if } x \in (m_{k-1}, m_k), k = 1, ..., N \\ (0, ..., 0, 1, 0, ..., 0) & \text{if } x = m_k, k = 0, 1, ..., N. \end{cases}$$

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Approximation of functions by epi-splines

lsc-fcns([l, u]) =lower semicontinuous functions (lsc) on [l, u] not $\neq \infty$

Adopt metric topology induced by the epi-distance

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Approximation of functions by epi-splines

lsc-fcns([I, u]) =lower semicontinuous functions (lsc) on [I, u] not $\neq \infty$

Adopt metric topology induced by the epi-distance

Need to allow for

- jumps (discontinuous densities)
- ▶ infinity (due to e^{-s})
- pointwise constraints (soft information)
- mixtures with probability mass functions
- subsequent maximization of densities (find modes)

Epi-distance

point-to-set distance $= d(x, S) = \inf_{y \in S} ||x - y||$ for $S \subset R^2$

Epi-distance

point-to-set distance
$$= d(x,S) = \inf_{y \in S} \|x-y\|$$
 for $S \subset R^2$

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epi-graph of
$$f = epi \ f = \{(x, \beta) \in \mathbb{R}^2 \mid f(x) \le \beta\}$$

Epi-distance

point-to-set distance
$$= d(x,S) = \inf_{y \in S} \|x-y\|$$
 for $S \subset R^2$

epi-graph of
$$f = \operatorname{epi} f = \{(x, \beta) \in \mathbb{R}^2 \mid f(x) \leq \beta\}$$

Epi-distance
$$= dl(f,g) = \int_0^\infty dl_\rho(f,g)e^{-\rho}d\rho$$

where for $\rho \geq \mathbf{0}$

$$dl_{
ho}(f,g) = \max_{\|x\| \leq
ho} |d(x, ext{epi } f) - d(x, ext{epi } g)|$$

Epi-distance (cont.)

 $dl_{
ho}(f,g) = \max_{\|x\| \leq
ho} |d(x, ext{epi } f) - d(x, ext{epi } g)|$



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Epi-distance (cont.)

$$dI_{
ho}(f,g) = \max_{\|x\| \le
ho} |d(x, ext{epi } f) - d(x, ext{epi } g)|$$



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Characterization of convergence

For any
$$\bar{
ho} \ge 0$$
:
 $dl(f^{
u}, f) \to 0 \iff dl_{
ho}(f^{
u}, f) \to 0 \text{ for all } \rho \ge \bar{
ho}$

(lsc-fcns([I, u]), dI) complete separable metric space

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Convergence of exponential epi-splines

exponential epi-splines = x-spl^p(m) = { $e^{-s} | s \in e$ -spl^p(m)} hypo-distance $dI_{hypo}(f,g) = dI(-f,-g)$

 $h^{\nu}, h^0 \in x$ -spl^p(m), $h^{\nu} = e^{-s^{\nu}} = e^{-\langle c(\cdot), r^{\nu} \rangle}$, $h^0 = e^{-s^0} = e^{-\langle c(\cdot), r^0 \rangle}$ Then, the following hold:

$$r^{
u}
ightarrow r^{0} \iff h^{
u}
ightarrow h^{0}$$
 uniformly on $[m_{0}, m_{N}]$
 $\implies dl_{\text{hypo}} (h^{
u}, h^{0})
ightarrow 0 \iff dl(s^{
u}, s^{0})
ightarrow 0$

Moreover, if h^{ν} , h^{0} are usc, then also

 $h^{
u}
ightarrow h^{0}$ uniformly on $[m_{0},m_{N}] \Longleftarrow dl_{\mathrm{hypo}} \ (h^{
u},h^{0})
ightarrow 0$

Properties of (exponential) epi-splines

If $\{m^
u\}_{
u=1}^\infty$ are refining meshes, then

► lsc
$$\{e-spl^p(m^\nu)\}_{\nu=1}^{\infty}$$
 dense in $lsc-fcns([I, u])$

• usc
$$\{x-\operatorname{spl}^p(m^\nu)\}_{\nu=1}^\infty$$
 dense in $\{e^{-s} \mid s \in \operatorname{lsc-fcns}([I, u])\}$

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most densities approximated to arbitrary accuracy

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Examples

- most densities approximated to arbitrary accuracy
- normal density represented by e-spl²(m) on [m₀, m_N] for any m

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Examples

- most densities approximated to arbitrary accuracy
- ▶ normal density represented by e-spl²(m) on [m₀, m_N] for any m
- exponential density represented by e-spl¹(m) on [m₀, m_N] for any choice of m with m₀ = 0

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Examples

- most densities approximated to arbitrary accuracy
- ▶ normal density represented by e-spl²(m) on [m₀, m_N] for any m
- exponential density represented by e-spl¹(m) on [m₀, m_N] for any choice of m with m₀ = 0

 lognormal and Pareto also exactly represented after transformation

Recall: maximize log-likelihood

$$h^{\nu} \in \operatorname{argmax}_{h} \frac{1}{\nu} \sum_{i=1}^{\nu} \log h(X^{i})$$

s.t.
$$\int_{-\infty}^{\infty} h(x) dx = 1$$
$$h \geq 0$$
$$h \in H^{\nu} \subset \mathcal{H}$$

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Now $h = e^{-\langle c(\cdot), r \rangle}$ and optimize over $r \in \mathbb{R}^{(p+2)N+1}$ instead

Resulting optimization problem

s.t.
$$\begin{split} \min_{r} \frac{1}{\nu} \sum_{i=1}^{\nu} \langle c(X^{i}), r \rangle \\ \int_{m_{0}}^{m_{N}} e^{-\langle c(x), r \rangle} dx &= 1 \\ r \in R^{\nu} &= \left\{ r \in R^{(p+2)N+1} \mid e^{-\langle c(\cdot), r \rangle} \in H^{\nu} \right\} \end{split}$$

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 R^{ν} often convex; \leq often replaces = \implies convex problem

Estimator unique under additional assumptions

Outline

Density estimation as a stochastic program

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- Epi-splines: pliable approximation tools
- Consistency and asymptotics
- Fusion of hard and soft information
- Numerical examples

Consistency

Kullback-Leibler projection of density h on $e-spl^{p}(m)$ is the set

$$\mathcal{P}_{p,m}(h) = \operatorname*{argmin}_{s \in \mathrm{e-spl}^p(m)} d_{KL}(h||e^{-s}) \text{ s.t. } \int_{m_0}^{m_N} e^{-s(x)} dx = 1$$

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 $\mathcal{P}_{p,m}^{S}(h) = \mathsf{KL} ext{-projection relative to } S \subset \operatorname{e-spl}^{p}(m)$

Consistency

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 $\mathcal{P}^{S}_{p,m}(h) = \mathsf{KL} ext{-projection relative to } S \subset \operatorname{e-spl}^{p}(m)$

$$\begin{aligned} P_{p,m}^{\nu} : \quad s^{\nu} \in \operatorname*{argmin}_{s \in S^{\nu}} \frac{1}{\nu} \sum_{i=1}^{\nu} s(X^{i}) \quad \text{s.t.} \quad \int_{m_{0}}^{m_{N}} e^{-s(x)} dx = 1 \\ S^{\nu} = \left\{ s \in \operatorname{e-spl}^{p}(m) \ \middle| \ e^{-s} \in H^{\nu} \right\} \end{aligned}$$

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Consistency (cont.)

True density $h^0 = e^{-s^0}$, with $s^0 = \langle c(\cdot), r^0 \rangle \in e\text{-spl}^p(m)$ Independent sample from h^0 $\{s^{\nu}\}_{\nu=1}^{\infty}$ sequence of optimal solutions of $P_{p,m}^{\nu}$, with $\{r^{\nu}\}_{\nu=1}^{\infty}$

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Consistency (cont.)

True density $h^0 = e^{-s^0}$, with $s^0 = \langle c(\cdot), r^0 \rangle \in e\text{-spl}^p(m)$ Independent sample from h^0 $\{s^{\nu}\}_{\nu=1}^{\infty}$ sequence of optimal solutions of $P_{p,m}^{\nu}$, with $\{r^{\nu}\}_{\nu=1}^{\infty}$

If $\lim R^{\nu}$ exists almost surely and is deterministic, then every accumulation point r^{∞} of $\{r^{\nu}\}_{\nu=1}^{\infty}$ satisfies

$$\langle c(\cdot), r^{\infty}
angle \in \mathcal{P}^{S^{\infty}}_{p,m}(h^0)$$
 almost surely,

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where $S^{\infty} = \{ s \in \text{e-spl}^{p}(m) \mid s = \langle c(\cdot), r \rangle, r \in \lim R^{\nu} \}$

Consistency (cont.)

Regardless of whether R^{ν} has a limit, if there exists $\{\hat{r}^{\nu}\}_{\nu=1}^{\infty}$, $\hat{r}^{\nu} \in R^{\nu}$, such that $\hat{r}^{\nu} \to r^{0}$ a.s., then a.s.: (i) $\langle c(\cdot), r^{\infty} \rangle \in \mathcal{P}_{p,m}(h^{0})$ (ii) $r_{\text{ess}}^{\infty} = r_{\text{ess}}^{0}$ ('essential' part: $r^{0} = (r_{\text{mesh}}^{0}, r_{\text{ess}}^{0}))$ (iii) If $r^{\nu} \to K r^{\infty}$ along a subsequence K, then $\langle c(\cdot), r^{\nu} \rangle \to K s^{0}$ and $e^{-\langle c(\cdot), r^{\nu} \rangle} \to K h^{0}$

uniformly on $[m_0, m_N]$, possibly except on m

Proof of consistency

- Let X^0 have density h^0 .
- Since X⁰ ∈ [m₀, m_N] almost surely, c(X⁰) is a random vector with finite moments

- Law of large number $(1/\nu) \sum_{i=1}^{\nu} c(X^i) \rightarrow E\{c(X^0)\}$ a.s.
- Epi-convergence of (effective) objective functions follow

Stability of Kullback-Leibler projection

Densities h^{ν} , h^{0} on [I, u] satisfy $dI_{\rm hypo}$ $(h^{\nu}, h^{0}) \rightarrow 0$

If r^{ν} is such that

$$\langle c(\cdot), r^{\nu} \rangle \in \mathcal{P}_{p,m}(h^{\nu}) \text{ for } m = \{m_k\}_{k=0}^N, m_0 = l, m_N = u,$$

then every accumulation point of $\{r^{\nu}\}_{\nu=1}^{\infty}$ is the epi-spline parameter of some $s^0 \in \mathcal{P}_{p,m}(h^0)$

Connections between modes of convergence

Densities $h^{\nu}, h^0 \in x$ -spl^p(m), with $h^{\nu} = e^{-\langle c(\cdot), r^{\nu} \rangle}, h^0 = e^{-\langle c(\cdot), r^0 \rangle}$ Then,

$$r^{\nu}
ightarrow r^{0} \Longrightarrow d_{\mathcal{KL}}(h^{0}||h^{\nu})
ightarrow 0 \Longleftrightarrow d_{\mathcal{KL}}(h^{\nu}||h^{0})
ightarrow 0 \Longrightarrow r^{\nu}_{\mathrm{ess}}
ightarrow r^{0}_{\mathrm{ess}}$$

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Asymptotic normality

True density $h^0 = e^{-s^0} \in \operatorname{x-spl}^p(m)$, $s^0 = \langle c(\cdot), r^0 \rangle$ r^0 in the interior of lim inf R^{ν} a.s. Independent sample from h^0 $\{s^{\nu}\}_{\nu=1}^{\infty}$ optimal solutions of $P_{\rho,m}^{\nu}$, with $\{r^{\nu}\}_{\nu=1}^{\infty}$, $h^{\nu} = e^{-\langle c(\cdot), r^{\nu} \rangle}$

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Then:

$$\begin{split} \nu^{1/2}(r_{\rm ess}^{\nu}-r_{\rm ess}^{0}) \rightarrow^{d} \mathcal{N}(0,\Sigma(r_{\rm ess}^{0})) \\ \nu^{1/2}(h^{\nu}(x)-h^{0}(x)) \rightarrow^{d} \mathcal{N}\left(0,\Sigma_{x}(r_{\rm ess}^{0}), \ x \in (m_{k-1},m_{k}), k=1,...,N \right. \\ \text{Moment estimator } \mu_{j}^{\nu} = \int_{m_{0}}^{m_{N}} x^{j} e^{-\langle c(x),r^{\nu} \rangle} dx \text{ satisfies} \end{split}$$

$$u^{1/2}(\mu_j^
u-\mu_j^0)
ightarrow^d\mathcal{N}(0,\langle w,\Sigma(r_{\mathrm{ess}}^0)w
angle)$$

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Outline

Density estimation as a stochastic program

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- Epi-splines: pliable approximation tools
- Consistency and asymptotics
- Fusion of hard and soft information
- Numerical examples

Formulation of soft information

Easy to ensure bounds on domain, continuity, smoothness, monotonicity



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Continuity

Epi-spline parameter

$$r = (s_0, ..., s_N, a_{1,0}, a_{1,1}, ..., a_{1,p}, a_{2,0}, a_{2,1}, ..., a_{2,p}, ..., a_{N,0}, a_{N,1}, ..., a_{N,p})$$

$$s_{k-1} = a_{k,0}, \ \ s_k = \sum_{i=0}^{p} a_{k,i} (m_k - m_{k-1})^i, \ \ k = 1, 2, ..., N$$

Kullback-Leibler constraint

Recall: Kullback-Leibler divergence from density h to density g

$$d_{\mathcal{K}L}(h||g) = \int_{-\infty}^{\infty} h(x) \log \frac{h(x)}{g(x)} dx$$

If $s \in e-spl^{p}(m)$ and r its epi-spline parameter, then

$$d_{\mathcal{K}L}(h||e^{-s}) = \left\langle \int_{m_0}^{m_N} c(x)h(x)dx, r \right\rangle + \int_{-\infty}^{\infty} (\log h(x))h(x)dx,$$

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So $\kappa_1 \leq d_{\mathcal{KL}}(h||e^{-s}) \leq \kappa_2$ are linear constraints

Outline

Density estimation as a stochastic program

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Example: 2-dof dynamical system



$$m_1\ddot{u}_1(t) + (k_1 + k_2)u_1(t) - k_2u_2(t) = p_o\sin vt$$

 $m_2\ddot{u}_2(t) - k_2u_1(t) + k_2u_2(t) = 0$

For choices of k_i , m_i , steady-state displacement at node 2:

$$u_2(t) = u_{2o} \sin vt$$
 with amplitude $u_{2o} = \frac{1}{(1 - v^2)(1 - v^2/4)}$

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Example: Density of response amplitude

V mix of beta densities gives density for amplitude X:



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Example: Density of response amplitude (cont.)

Sample size 100; continuously differentiable; "unimodal" tails



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Gradient information

Gradient information for bijective $G : R \to R$ Recall: If X = G(V), then

$$h_X(x) = h_V(G^{-1}(x))/|G'(G^{-1}(x))|$$

Present context without a bijection and data $x^i = G(v^i), G'(v^i)$:

$$egin{aligned} h^
u(x^i) &= e^{-\langle c(x^i), r
angle} \geq rac{h_V(v^i)}{|G'(v^i)|} \ \langle c(x^i), r
angle \leq -\log rac{h_V(v^i)}{|G'(v^i)|} \end{aligned}$$

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Value of pdf bounded from below at x^i

Back to example

Sample size 20; gradient information



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Example: Uniform mixture density sample size 1000; lsc



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Example: Uniform mixture density (cont.) sample size just 100; lsc



Summary

- Density estimation problems as stochastic programs
- Exponential epi-splines offer a tractable class of density estimators
- Incorporate soft information by means of constraints
- Extensions to response surface, regression curve, multivariate density estimation, and many other curve fitting problems

References

- Royset & Wets, "Nonparametric density estimation via exponential epi-splines: fusion of soft and hard information"
- Royset & Wets, "Epi-splines and exponential epi-splines: Pliable approximation tools"
- Singham, Royset, & Wets, "Density estimation of simulation output using exponential epi-splines," *Proc. WSC*, 2013
- Royset, Sukumar, & Wets "Uncertainty quantification using exponential epi-splines," *Proc. ICOSSAR*, 2013

- More examples: https://www.math.ucdavis.edu/~prop01/
- Matlab implementations available